Money may buy happiness, but often so little that it doesn’t matter


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Money may buy happiness, but often so little that it doesn’t matter

Whether money buys happiness or not is a question of enduring individual and societal interest that has justifiably attracted considerable attention from researchers across the social sciences (Clark, Frijters, & Shields, 2008; Kahneman & Deaton, 2010). Consistently, research points toward there being a weak relationship (Lucas & Dyrenforth, 2006), leading many to conclude that to improve people's lives we will have to go beyond focusing on money (Diener & Seligman, 2004). However, one interesting research stream suggests that the weak relationship between money and happiness arises because people do not spend their money wisely (Dunn, Gilbert, & Wilson, 2011). The implication being that more money would translate into greater happiness if people spent it “right”; for example, on experiences rather than possessions (Van Boven & Gilovich, 2003), or on others rather than themselves (Dunn, Aknin, & Norton, 2008). In line with this interesting and influential research stream, Matz, Gladstone, and Stillwell (2016) drew on a rich dataset of more than 76,000 bank transactions (N=625 bank customers recruited from 150,000 invited to participate) to examine whether individuals who spend on goods that match their personality are more satisfied than those who do not. They then followed up with a study showing that students (N=79) randomized into spending in line with their personality had higher levels of happiness. Their finding is important in that it further highlights the potential role of personality in moderating the influence of income and spending on well-being (Boyce, Wood, & Ferguson, 2016; Soto & Luhmann, 2013). However, the extent to which Matz et al. (2016) provided evidence that spending more money can buy happiness if it is spent “right”, as implied by the article, is overstated.

In Study 1 Matz et al. showed that people spent more money on personality-matched products (Table 2) and that people who spent on personality-matched products have marginally higher life satisfaction (Table 3). However, it is impossible to link the two analyses because the second stage did not include the actual amount spent on personality-matched products. It therefore cannot be inferred that if spent “right” there would be a stronger relationship between levels of consumption and well-being or that the findings offer a “contrast to decades of research reporting surprisingly weak relationships between consumption and happiness” (Matz et al., 2016, p. 715). To evaluate whether support existed for this key idea it would have been necessary to examine whether the relationship between total spending and life-satisfaction was moderated by the strength of the match between the buyers’ personality and their purchases. Crucially, the study did not do this.
Instead, the authors pointed to potential gains in life-satisfaction associated with personality-matched purchases irrespective of the amount spent. To show this the authors calculated the difference between the personality z-scores (e.g. openness = O, neuroticism = N) of the participant (i) and z-scores of the personality profile of the participants shopping basket (b). They then rescaled this measure so that moving from low to high scores indicated a greater match between participants and basket personality scores (see equation 1):

\[
n_{i,b} = \text{mean} - \sqrt{(z(O_i) - z(O_b))^2 + \cdots + (z(N_i) - z(N_b))^2}
\]

However, it’s not clear whether this somewhat complex matching variable captures meaningful variation in personality-matched purchasing. For example, participant and basket personality scores appear to be very weakly correlated (average \( r = .05 \), as shown in their supplementary Table S2) suggesting that few people tend to make purchases that closely match their personality. This led us to speculate that the only statistically significant link between personality-matched spending and life satisfaction presented in the paper (Model 1, Table 3) could have been driven by participant personality, which is known to relate to well-being (Diener & Lucas, 1999). Although we do not have access to the sensitive banking data used by Matz et al., we could generate a basket-participant match variable where participant and basket personality scores were similarly uncorrelated. We did this by matching randomly generated personality profiles for a participant’s shopping basket to real personality data from participants in the English Longitudinal Study of Ageing (ELSA). We repeated this randomization 1000 times in a sample of equivalent size (\( N=625 \)) to Matz et al. to gauge the likelihood of detecting a spurious effect where one would not be expected. The results of our analyses are shown in Table 1 where we observe that greater \( n_{i,b} \) positively predicts life satisfaction (\( \beta = .12 \)) as in Matz et al. (Model 1, Table 3). This occurred in 80.3% of our replications (where \( p < .05 \)) and 100% of replications when the full ELSA sample (\( N = 7,990 \)) was used. Whilst this analysis does not disprove the results of Matz et al, it highlights that it is difficult to precisely gauge what the matching variable used in the study is assessing given a similar pattern of results can be found when personality scores for purchases are randomly generated.

Nevertheless, even if the results are taken at face value they appear to show that attempting to spend one’s way to happiness would “buy” so little well-being as to be largely irrelevant to people’s lives. Standardized effect sizes are not provided, yet the R-squared values in their Table 3 suggest that the extent to which an individual spends in line with their personality explains less than 1% of the variance in life satisfaction. Income typically
Table 1

Average Results Based on 1000 Random Repetitions of Multiple Linear Regression Analyses

Predicting Life Satisfaction Using Participant Personality Data From a Random Sample of the English Longitudinal Study of Ageing (N = 625) Matched to Randomly Generated Spending/Shopping Basket Personality Variables

<table>
<thead>
<tr>
<th>VARIABLES</th>
<th>(1) Dependent variable: Life Satisfaction (standardized)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Basket-participant match</td>
<td>0.11** (0.04)</td>
</tr>
<tr>
<td>Household income (log)</td>
<td>0.25** (0.06)</td>
</tr>
<tr>
<td>Gender</td>
<td>-0.05 (0.08)</td>
</tr>
<tr>
<td>Age</td>
<td>0.00 (0.00)</td>
</tr>
<tr>
<td>Percentage of random draws in which basket-participant match is significant (N = 625)</td>
<td>80.3%</td>
</tr>
<tr>
<td>Percentage of random draws in which basket-participant match is significant in the full sample (N = 7,990)</td>
<td>100%</td>
</tr>
</tbody>
</table>

Note: We randomly selected a sample of 625 of a possible 7,990 who answered all questions on life satisfaction, personality, and household income, in wave 5 of the English Longitudinal Study of Ageing. We then randomly generated personality z-scores for each participants “basket personality”. We calculated the basket-participant match using the Euclidean distance following Matz et al. as follows:

\[
(basket - participant match_{i,b} = mean - \left( \sqrt{ (z(O_1) - z(O_b))^2 + \cdots + (z(N_1) - z(N_b))^2 } \right), \text{ where } i \text{ represents the participant z-scored personality and } b \text{ represents the individuals (randomly generated) z-scored shopping basket personality characteristics. We then repeated this 1000 times with the results representing the average across the 1000 repetitions. We also carried out this analysis on the full ELSA sample. Standard errors in parentheses; ** p<0.01, * p<0.05.}
explains at best 4% (Lucas & Dyrenforth, 2006) and given that consumption expenditures are as important to life satisfaction as income (Headey, Muffels, & Wooden, 2007) their results do not support the implication that the relationship is of meaningful magnitude. Thus, whilst money may buy happiness through consumption the conclusion should be, unless clearly demonstrated otherwise, that it buys so little it doesn’t matter.

In contrast to this work, psychologists have identified several areas where money may matter to well-being. For example, income can help alleviate the psychological distress from poverty (Weich & Lewis, 1998) and indebtedness (Gathergood, 2012). Further, low income rank appears to adversely influence well-being (Boyce, Brown, & Moore, 2010) as does losing income (Boyce, Wood, Banks, Clark, & Brown, 2013), particularly for those who value economic outcomes greatly (Boyce et al., 2016). However, even the magnitude of these robust effects tends to be dwarfed by the importance (as indexed by standardized effect sizes) of other factors such as relationships (Powdthavee, 2008), stable employment (McKee-Ryan, Song, Wanberg, & Kinicki, 2005), and mental and physical health (Layard, Clark, Cornaglia, Powdthavee, & Vernoit, 2014). Personality itself, irrespective of any interactive effect, has long been regarded as one of the strongest predictors of happiness (Diener & Lucas, 1999). Indeed Matz et al. showed in their Table 3 that the main effects of neuroticism and extraversion are both far more important than any other variable in the study.

Given that Matz et al. (2016) showed the relative unimportance of consumption in providing people with greater happiness, regardless of how it is spent, it is a concern when it is implied that a consumption-based society can be a key route to greater happiness for all. Materialism is associated with less happiness (Dittmar, Bond, Hurst, & Kasser, 2014) and there is a danger that if academics encourage people to pursue consumption with the hope of obtaining greater happiness, then individuals and policy makers may sacrifice pursuing the very things that are the most important to it.
References


